

# Hypothesis Testing

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# Basics of Hypothesis Testing

# What is a Hypothesis?

One situation among a set of possible situations

## Example (Radar)

EM waves are transmitted and the reflections observed.

**Null Hypothesis** Plane absent

**Alternative Hypothesis** Plane present

For a given set of observations, either hypothesis may be true.

# What is Hypothesis Testing?

- A statistical framework for deciding which hypothesis is true
- Under each hypothesis the observations are assumed to have a known distribution
- Consider the case of two hypotheses (binary hypothesis testing)

$$\begin{aligned}H_0 &: \mathbf{Y} \sim P_0 \\H_1 &: \mathbf{Y} \sim P_1\end{aligned}$$

$\mathbf{Y}$  is the random observation vector belonging to observation set  $\Gamma \subseteq \mathbb{R}^n$  for  $n \in \mathbb{N}$

- The hypotheses are assumed to occur with given prior probabilities

$$\begin{aligned}\Pr(H_0 \text{ is true}) &= \pi_0 \\ \Pr(H_1 \text{ is true}) &= \pi_1\end{aligned}$$

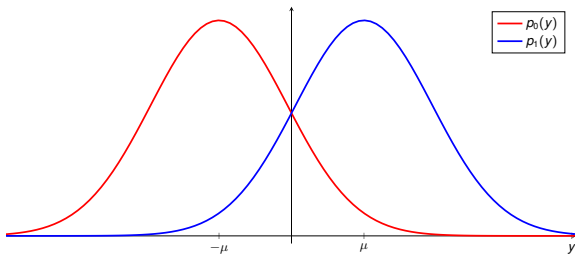
where  $\pi_0 + \pi_1 = 1$ .

# Location Testing with Gaussian Error

- Let observation set  $\Gamma = \mathbb{R}$  and  $\mu > 0$

$$H_0 : Y \sim N(-\mu, \sigma^2)$$

$$H_1 : Y \sim N(\mu, \sigma^2)$$



- Any point in  $\Gamma$  can be generated under both  $H_0$  and  $H_1$
- What is a **good decision rule** for this hypothesis testing problem which takes the prior probabilities into account?

# What is a Decision Rule?

- A decision rule for binary hypothesis testing is a partition of  $\Gamma$  into  $\Gamma_0$  and  $\Gamma_1$  such that

$$\delta(\mathbf{y}) = \begin{cases} 0 & \text{if } \mathbf{y} \in \Gamma_0 \\ 1 & \text{if } \mathbf{y} \in \Gamma_1 \end{cases}$$

We decide  $H_i$  is true when  $\delta(\mathbf{y}) = i$  for  $i \in \{0, 1\}$

- For the location testing with Gaussian error problem, one possible decision rule is

$$\Gamma_0 = (-\infty, 0]$$

$$\Gamma_1 = (0, \infty)$$

and another possible decision rule is

$$\Gamma_0 = (-\infty, -100) \cup (-50, 0)$$

$$\Gamma_1 = [-100, -50] \cup [0, \infty)$$

- Given that partitions of the observation set define decision rules, what is the optimal partition?

# Which is the Optimal Decision Rule?

- The optimal decision rule minimizes the probability of decision error
- For the binary hypothesis testing problem of  $H_0$  versus  $H_1$ , the conditional decision error probability given  $H_i$  is true is

$$\begin{aligned}P_{e|i} &= \Pr[\text{Deciding } H_{1-i} \text{ is true} | H_i \text{ is true}] \\ &= \Pr[Y \in \Gamma_{1-i} | H_i] \\ &= 1 - \Pr[Y \in \Gamma_i | H_i] \\ &= 1 - P_{c|i}\end{aligned}$$

- Probability of decision error is

$$P_e = \pi_0 P_{e|0} + \pi_1 P_{e|1}$$

- Probability of correct decision is

$$P_c = \pi_0 P_{c|0} + \pi_1 P_{c|1} = 1 - P_e$$

# Which is the Optimal Decision Rule?

- Maximizing the probability of correct decision will minimize probability of decision error
- Probability of correct decision is

$$\begin{aligned} P_c &= \pi_0 P_{c|0} + \pi_1 P_{c|1} \\ &= \pi_0 \int_{y \in \Gamma_0} p_0(y) dy + \pi_1 \int_{y \in \Gamma_1} p_1(y) dy \end{aligned}$$

- If a point  $y$  in  $\Gamma$  belongs to  $\Gamma_i$ , its contribution to  $P_c$  is proportional to  $\pi_i p_i(y)$
- To maximize  $P_c$ , we choose the partition  $\{\Gamma_0, \Gamma_1\}$  as

$$\begin{aligned} \Gamma_0 &= \{y \in \Gamma | \pi_0 p_0(y) \geq \pi_1 p_1(y)\} \\ \Gamma_1 &= \{y \in \Gamma | \pi_1 p_1(y) > \pi_0 p_0(y)\} \end{aligned}$$

- The points  $y$  for which  $\pi_0 p_0(y) = \pi_1 p_1(y)$  can be in either  $\Gamma_0$  and  $\Gamma_1$  (the optimal decision rule is not unique)



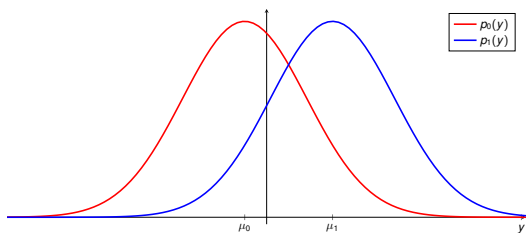
# Location Testing with Gaussian Error

- Let  $\mu_1 > \mu_0$  and  $\pi_0 = \pi_1 = \frac{1}{2}$

$$H_0 : Y = \mu_0 + Z$$

$$H_1 : Y = \mu_1 + Z$$

where  $Z \sim N(0, \sigma^2)$



$$p_0(y) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y-\mu_0)^2}{2\sigma^2}}$$

$$p_1(y) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y-\mu_1)^2}{2\sigma^2}}$$

# Location Testing with Gaussian Error

- Optimal decision rule is given by the partition  $\{\Gamma_0, \Gamma_1\}$

$$\Gamma_0 = \{y \in \Gamma \mid \pi_0 p_0(y) \geq \pi_1 p_1(y)\}$$

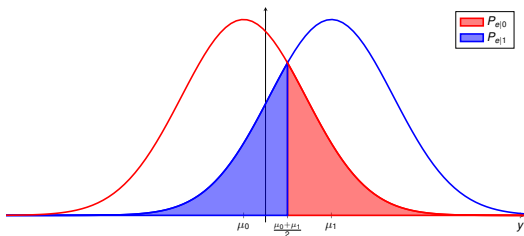
$$\Gamma_1 = \{y \in \Gamma \mid \pi_1 p_1(y) > \pi_0 p_0(y)\}$$

- For  $\pi_0 = \pi_1 = \frac{1}{2}$

$$\Gamma_0 = \left\{ y \in \Gamma \mid y \leq \frac{\mu_1 + \mu_0}{2} \right\}$$

$$\Gamma_1 = \left\{ y \in \Gamma \mid y > \frac{\mu_1 + \mu_0}{2} \right\}$$

# Location Testing with Gaussian Error



$$P_{e|0} = \Pr \left[ Y > \frac{\mu_0 + \mu_1}{2} \mid H_0 \right] = Q \left( \frac{\mu_1 - \mu_0}{2\sigma} \right)$$

$$P_{e|1} = \Pr \left[ Y \leq \frac{\mu_0 + \mu_1}{2} \mid H_1 \right] = \Phi \left( \frac{\mu_0 - \mu_1}{2\sigma} \right) = Q \left( \frac{\mu_1 - \mu_0}{2\sigma} \right)$$

$$P_e = \pi_0 P_{e|0} + \pi_1 P_{e|1} = Q \left( \frac{\mu_1 - \mu_0}{2\sigma} \right)$$

This  $P_e$  is for  $\pi_0 = \pi_1 = \frac{1}{2}$

# Location Testing with Gaussian Error

- Suppose  $\pi_0 \neq \pi_1$
- Optimal decision rule is still given by the partition  $\{\Gamma_0, \Gamma_1\}$

$$\Gamma_0 = \{y \in \Gamma \mid \pi_0 p_0(y) \geq \pi_1 p_1(y)\}$$

$$\Gamma_1 = \{y \in \Gamma \mid \pi_1 p_1(y) > \pi_0 p_0(y)\}$$

- The partitions specialized to this problem are

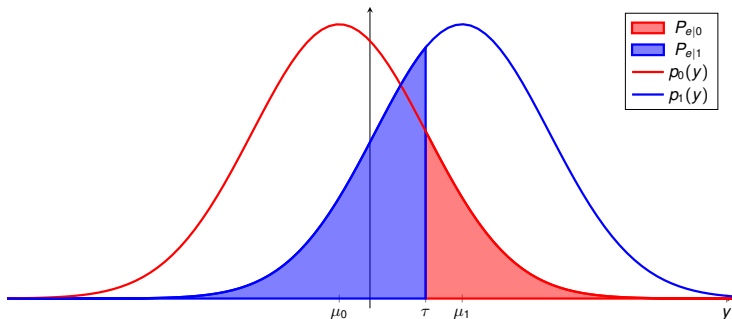
$$\Gamma_0 = \left\{ y \in \Gamma \mid y \leq \frac{\mu_1 + \mu_0}{2} + \frac{\sigma^2}{(\mu_1 - \mu_0)} \log \frac{\pi_0}{\pi_1} \right\}$$

$$\Gamma_1 = \left\{ y \in \Gamma \mid y > \frac{\mu_1 + \mu_0}{2} + \frac{\sigma^2}{(\mu_1 - \mu_0)} \log \frac{\pi_0}{\pi_1} \right\}$$

# Location Testing with Gaussian Error

Suppose  $\pi_0 = 0.6$  and  $\pi_1 = 0.4$

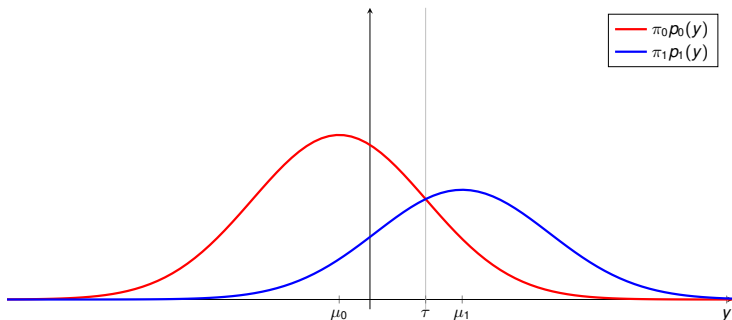
$$\tau = \frac{\mu_1 + \mu_0}{2} + \frac{\sigma^2}{(\mu_1 - \mu_0)} \log \frac{\pi_0}{\pi_1} = \frac{\mu_1 + \mu_0}{2} + \frac{0.4054\sigma^2}{(\mu_1 - \mu_0)}$$



# Location Testing with Gaussian Error

Suppose  $\pi_0 = 0.6$  and  $\pi_1 = 0.4$

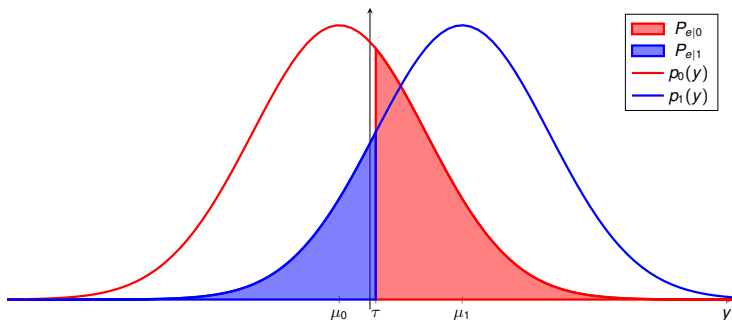
$$\tau = \frac{\mu_1 + \mu_0}{2} + \frac{\sigma^2}{(\mu_1 - \mu_0)} \log \frac{\pi_0}{\pi_1} = \frac{\mu_1 + \mu_0}{2} + \frac{0.4054\sigma^2}{(\mu_1 - \mu_0)}$$



# Location Testing with Gaussian Error

Suppose  $\pi_0 = 0.4$  and  $\pi_1 = 0.6$

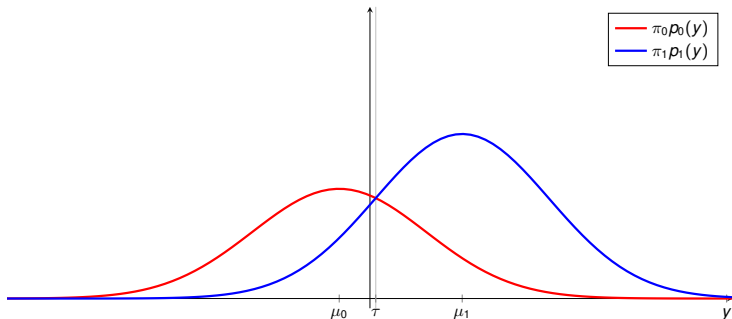
$$\tau = \frac{\mu_1 + \mu_0}{2} + \frac{\sigma^2}{(\mu_1 - \mu_0)} \log \frac{\pi_0}{\pi_1} = \frac{\mu_1 + \mu_0}{2} - \frac{0.4054\sigma^2}{(\mu_1 - \mu_0)}$$



# Location Testing with Gaussian Error

Suppose  $\pi_0 = 0.4$  and  $\pi_1 = 0.6$

$$\tau = \frac{\mu_1 + \mu_0}{2} + \frac{\sigma^2}{(\mu_1 - \mu_0)} \log \frac{\pi_0}{\pi_1} = \frac{\mu_1 + \mu_0}{2} - \frac{0.4054\sigma^2}{(\mu_1 - \mu_0)}$$





# $M$ -ary Hypothesis Testing

- $M$  hypotheses with prior probabilities  $\pi_i, i = 1, \dots, M$

$$H_1 : \mathbf{Y} \sim P_1$$

$$H_2 : \mathbf{Y} \sim P_2$$

$$\vdots$$

$$H_M : \mathbf{Y} \sim P_M$$

- A decision rule for  $M$ -ary hypothesis testing is a partition of  $\Gamma$  into  $M$  disjoint regions  $\{\Gamma_i | i = 1, \dots, M\}$  such that

$$\delta(\mathbf{y}) = i \text{ if } \mathbf{y} \in \Gamma_i$$

We decide  $H_i$  is true when  $\delta(\mathbf{y}) = i$  for  $i \in \{1, \dots, M\}$

- Minimum probability of error rule is

$$\delta_{\text{MPE}}(\mathbf{y}) = \arg \max_{1 \leq i \leq M} \pi_i p_i(\mathbf{y})$$

# Maximum A Posteriori Decision Rule

- The a posteriori probability of  $H_i$  being true given observation  $\mathbf{y}$  is

$$P \left[ H_i \text{ is true} \middle| \mathbf{y} \right] = \frac{\pi_i p_i(\mathbf{y})}{p(\mathbf{y})}$$

- The MAP decision rule is given by

$$\delta_{\text{MAP}}(\mathbf{y}) = \arg \max_{1 \leq i \leq M} P \left[ H_i \text{ is true} \middle| \mathbf{y} \right] = \delta_{\text{MPE}}(\mathbf{y})$$

MAP decision rule = MPE decision rule

# Maximum Likelihood Decision Rule

- The ML decision rule is given by

$$\delta_{\text{ML}}(\mathbf{y}) = \arg \max_{1 \leq i \leq M} p_i(\mathbf{y})$$

- If the  $M$  hypotheses are equally likely,  $\pi_i = \frac{1}{M}$
- The MPE decision rule is then given by

$$\delta_{\text{MPE}}(\mathbf{y}) = \arg \max_{1 \leq i \leq M} \pi_i p_i(\mathbf{y}) = \delta_{\text{ML}}(\mathbf{y})$$

For equal priors, ML decision rule = MPE decision rule

## Irrelevant Statistics

# Irrelevant Statistics

- In this context, the term statistic means an observation
- For a given hypothesis testing problem, all the observations may not be useful

## Example (Irrelevant Statistic)

$$\mathbf{Y} = [Y_1 \quad Y_2]$$

$$H_1 : Y_1 = A + N_1, \quad Y_2 = N_2$$

$$H_0 : Y_1 = N_1, \quad Y_2 = N_2$$

where  $A > 0$ ,  $N_1 \sim N(0, \sigma^2)$ ,  $N_2 \sim N(0, \sigma^2)$ .

- If  $N_1$  and  $N_2$  are independent,  $Y_2$  is irrelevant.
- If  $N_1$  and  $N_2$  are correlated,  $Y_2$  is relevant.
- Need a method to recognize irrelevant components of the observations

# Characterizing an Irrelevant Statistic

## Theorem

For  $M$ -ary hypothesis testing using an observation  $\mathbf{Y} = [\mathbf{Y}_1 \quad \mathbf{Y}_2]$ , the statistic  $\mathbf{Y}_2$  is irrelevant if the conditional distribution of  $\mathbf{Y}_2$ , given  $\mathbf{Y}_1$  and  $H_i$ , is independent of  $i$ . In terms of densities, the condition for irrelevance is

$$p(\mathbf{y}_2|\mathbf{y}_1, H_i) = p(\mathbf{y}_2|\mathbf{y}_1) \quad \forall i.$$

## Proof

$$\delta_{\text{MPE}}(\mathbf{y}) = \arg \max_{1 \leq i \leq M} \pi_i p_i(\mathbf{y}) = \arg \max_{1 \leq i \leq M} \pi_i p(\mathbf{y}|H_i)$$

$$\begin{aligned} p(\mathbf{y}|H_i) &= p(\mathbf{y}_1, \mathbf{y}_2|H_i) = p(\mathbf{y}_2|\mathbf{y}_1, H_i)p(\mathbf{y}_1|H_i) \\ &= p(\mathbf{y}_2|\mathbf{y}_1)p(\mathbf{y}_1|H_i) \end{aligned}$$

$$\delta_{\text{MPE}}(\mathbf{y}) = \arg \max_{1 \leq i \leq M} \pi_i p(\mathbf{y}_2|\mathbf{y}_1)p(\mathbf{y}_1|H_i) = \arg \max_{1 \leq i \leq M} \pi_i p(\mathbf{y}_1|H_i)$$

# Example of an Irrelevant Statistic

## Example (Independent Noise)

$$\mathbf{Y} = [Y_1 \quad Y_2]$$

$$H_1 : Y_1 = A + N_1, \quad Y_2 = N_2$$

$$H_0 : Y_1 = N_1, \quad Y_2 = N_2$$

where  $A > 0$ ,  $N_1 \sim N(0, \sigma^2)$ ,  $N_2 \sim N(0, \sigma^2)$ ,  $N_1 \perp N_2$ .

$$\rho(\mathbf{y}_2 | \mathbf{y}_1, H_0) = \rho(\mathbf{y}_2)$$

$$\rho(\mathbf{y}_2 | \mathbf{y}_1, H_1) = \rho(\mathbf{y}_2)$$

# Example of a Relevant Statistic

## Example (Correlated Noise)

$$\mathbf{Y} = [Y_1 \quad Y_2]^T$$

$$H_1 : Y_1 = A + N_1, \quad Y_2 = N_2$$

$$H_0 : Y_1 = N_1, \quad Y_2 = N_2$$

where  $A > 0$ ,  $N_1 \sim N(0, \sigma^2)$ ,  $N_2 \sim N(0, \sigma^2)$ ,  $\mathbf{C}_Y = \sigma^2 \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix}$

$$p(y_2|y_1, H_0) = \frac{1}{\sqrt{2\pi(1-\rho^2)\sigma^2}} e^{-\frac{(y_2 - \rho y_1)^2}{2(1-\rho^2)\sigma^2}},$$

$$p(y_2|y_1, H_1) = \frac{1}{\sqrt{2\pi(1-\rho^2)\sigma^2}} e^{-\frac{[y_2 - \rho(y_1 - A)]^2}{2(1-\rho^2)\sigma^2}}$$



Thanks for your attention