Optimal Receiver for the AWGN Channel

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Additive White Gaussian Noise Channel



$$y(t) = s(t) + n(t)$$

- s(t) Transmitted Signal
- y(t) Received Signal
- n(t) White Gaussian Noise

$$S_n(f)=\frac{N_0}{2}=\sigma^2$$

$$R_n(\tau) = \sigma^2 \delta(\tau)$$

M-ary Signaling in AWGN Channel

- One of *M* continuous-time signals $s_1(t), \ldots, s_M(t)$ is sent
- The received signal is the transmitted signal corrupted by AWGN
- *M* hypotheses with prior probabilities π_i , i = 1, ..., M

$$\begin{array}{rcl} H_1 & : & y(t) = s_1(t) + n(t) \\ H_2 & : & y(t) = s_2(t) + n(t) \\ \vdots & & \vdots \\ H_M & : & y(t) = s_M(t) + n(t) \end{array}$$

- The model implicitly assumes that
 - the delay has been estimated and
 - there is no attenuation or other distortion.
- Random variables are easier to handle than random processes
- We derive an equivalent *M*-ary hypothesis testing problem involving only random vectors
- Two digressions follow
 - Gaussian random processes
 - Signal space representation

Gaussian Random Processes

Gaussian Random Process

Definition

A random process X(t) is Gaussian if its samples $X(t_1), \ldots, X(t_n)$ are jointly Gaussian for any $n \in \mathbb{N}$ and distinct sample locations t_1, t_2, \ldots, t_n .

Let $\mathbf{X} = \begin{bmatrix} X(t_1) & \cdots & X(t_n) \end{bmatrix}^T$ be the vector of samples. The joint density is given by

$$\rho(\mathbf{x}) = \frac{1}{\sqrt{(2\pi)^n \det(\mathbf{C})}} \exp\left(-\frac{1}{2}(\mathbf{x} - \mathbf{m})^T \mathbf{C}^{-1}(\mathbf{x} - \mathbf{m})\right)$$

where

$$\mathbf{m} = \boldsymbol{E}[\mathbf{X}], \ \mathbf{C} = \boldsymbol{E}\left[(\mathbf{X} - \mathbf{m})(\mathbf{X} - \mathbf{m})^{T}\right]$$

Properties

- The mean and autocorrelation functions completely characterize a Gaussian random process.
- Wide-sense stationary Gaussian processes are strictly stationary.

White Gaussian Noise

Definition

A zero mean WSS Gaussian random process with power spectral density

$$S_n(f)=\frac{N_0}{2}$$

 $\frac{N_0}{2}$ is termed the two-sided PSD and has units Watts per Hertz.

Remarks

- Autocorrelation function $R_n(\tau) = \frac{N_0}{2}\delta(\tau)$
- Infinite Power! Ideal model of Gaussian noise occupying more bandwidth than the signals of interest.

White Gaussian Noise through Correlators

Consider the output of a correlator with WGN input

$$Z = \int_{-\infty}^{\infty} n(t)u(t) \, dt = \langle n, u \rangle$$

where u(t) is a deterministic finite-energy real signal

- Z is a Gaussian random variable
- The mean of Z is

$$E[Z] = \int_{-\infty}^{\infty} E[n(t)] u(t) dt = 0$$

• The variance of Z is

$$\operatorname{var}(Z) = E\left[\left(\langle n, u \rangle\right)^{2}\right] = E\left[\int_{-\infty}^{\infty} n(t)u(t) \, dt \int_{-\infty}^{\infty} n(s)u(s) \, ds\right]$$
$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} u(t)u(s)E\left[n(t)n(s)\right] \, dt \, ds$$
$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} u(t)u(s)\frac{N_{0}}{2}\delta(t-s) \, dt \, ds$$
$$= \frac{N_{0}}{2} \int_{-\infty}^{\infty} u^{2}(t) \, dt = \frac{N_{0}}{2} \|u\|^{2}$$

White Gaussian Noise through Correlators

Proposition

Let $u_1(t)$ and $u_2(t)$ be finite-energy real signals and let n(t) be WGN with PSD $S_n(t) = \frac{N_0}{2}$. Then $\langle n, u_1 \rangle$ and $\langle n, u_2 \rangle$ are jointly Gaussian with covariance

$$\operatorname{cov}(\langle n, u_1 \rangle, \langle n, u_2 \rangle) = \frac{N_0}{2} \langle u_1, u_2 \rangle.$$

Proof

To prove that $\langle n, u_1 \rangle$ and $\langle n, u_2 \rangle$ are jointly Gaussian, consider a linear combination $a \langle n, u_1 \rangle + b \langle n, u_2 \rangle$

$$a\langle n, u_1 \rangle + b\langle n, u_2 \rangle = \int_{-\infty}^{\infty} n(t) [au_1(t) + bu_2(t)] dt.$$

This is the result of passing n(t) through a correlator. So it is a Gaussian random variable.

White Gaussian Noise through Correlators Proof (continued)

$$\operatorname{cov}\left(\langle n, u_{1} \rangle, \langle n, u_{2} \rangle\right) = E\left[\langle n, u_{1} \rangle \langle n, u_{2} \rangle\right]$$

$$= E\left[\int_{-\infty}^{\infty} n(t)u_{1}(t) dt \int_{-\infty}^{\infty} n(s)u_{2}(s) ds\right]$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} u_{1}(t)u_{2}(s)E\left[n(t)n(s)\right] dt ds$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} u_{1}(t)u_{2}(s)\frac{N_{0}}{2}\delta(t-s) dt ds$$

$$= \frac{N_{0}}{2} \int_{-\infty}^{\infty} u_{1}(t)u_{2}(t) dt$$

$$= \frac{N_{0}}{2} \langle u_{1}, u_{2} \rangle$$

If $u_1(t)$ and $u_2(t)$ are orthogonal, $\langle n, u_1 \rangle$ and $\langle n, u_2 \rangle$ are independent.

Signal Space Representation

Signal Space Representation of Waveforms

Given M finite energy waveforms, construct an orthonormal basis

$$s_1(t), \dots, s_M(t) \to \underbrace{\phi_1(t), \dots, \phi_N(t)}_{\text{Orthonormal basis}}$$

$$\langle \phi_i, \phi_j \rangle = \int_{-\infty}^{\infty} \phi_i(t) \phi_j^*(t) dt = \begin{cases} 1 & \text{if } i = j \\ 0 & \text{otherwise} \end{cases}$$

Each s_i(t) is a linear combination of the basis vectors

$$s_i(t) = \sum_{n=1}^N s_{i,n} \phi_n(t), \quad i = 1, \dots, M$$

- $s_i(t)$ is represented by the vector $\mathbf{s}_i = \begin{bmatrix} s_{i,1} & \cdots & s_{i,N} \end{bmatrix}^T$
- The set {s_i : 1 ≤ i ≤ M} is called the signal space representation or constellation

Constellation Point to Waveform



Waveform to Constellation Point



Gram-Schmidt Orthogonalization Procedure

- Algorithm for calculating orthonormal basis for s₁(t),..., s_M(t)
- Consider M = 1

$$\phi_1(t) = \frac{s_1(t)}{\|s_1\|}$$

where $\|\boldsymbol{s}_1\|^2 = \langle \boldsymbol{s}_1, \boldsymbol{s}_1 \rangle$

Consider M = 2

$$\phi_1(t) = \frac{s_1(t)}{\|s_1\|}, \quad \phi_2(t) = \frac{\gamma(t)}{\|\gamma\|}$$

where $\gamma(t) = s_2(t) - \langle s_2, \phi_1 \rangle \phi_1(t)$

Consider M = 3

$$\phi_1(t) = \frac{\mathbf{s}_1(t)}{\|\mathbf{s}_1\|}, \quad \phi_2(t) = \frac{\gamma_1(t)}{\|\gamma_1\|}, \quad \phi_3(t) = \frac{\gamma_2(t)}{\|\gamma_2\|}$$

where

$$\begin{array}{rcl} \gamma_1(t) &=& s_2(t) - \langle s_2, \phi_1 \rangle \phi_1(t) \\ \gamma_2(t) &=& s_3(t) - \langle s_3, \phi_1 \rangle \phi_1(t) - \langle s_3, \phi_2 \rangle \phi_2(t) \end{array}$$

Gram-Schmidt Orthogonalization Procedure

• In general, given $s_1(t), \ldots, s_M(t)$ the *k*th basis function is

$$\phi_k(t) = \frac{\gamma_k(t)}{\|\gamma_k\|}$$

where

$$\gamma_k(t) = s_k(t) - \sum_{i=1}^{k-1} \langle s_k, \phi_i \rangle \phi_i(t)$$

is not the zero function

If γ_k(t) is zero, s_k(t) is a linear combination of φ₁(t),..., φ_{k-1}(t). It does not contribute to the basis.

Gram-Schmidt Procedure Example



Gram-Schmidt Procedure Example





Properties of Signal Space Representation

Energy

$$E_m = \int_{-\infty}^{\infty} |s_m(t)|^2 dt = \sum_{n=1}^{N} |s_{m,n}|^2 = ||\mathbf{s}_m||^2$$

Inner product

 $\langle s_i(t), s_j(t) \rangle = \langle \mathbf{s}_i, \mathbf{s}_j \rangle$

Optimal Receiver for the AWGN Channel

Restriction to Signal Space is Optimal

Theorem

For the M-ary hypothesis testing given by

$$H_1 : y(t) = s_1(t) + n(t)$$

$$\vdots \qquad \vdots$$

$$H_M : y(t) = s_M(t) + n(t)$$

there is no loss in detection performance by using the optimal decision rule for the following M-ary hypothesis testing problem

$$H_1 : \mathbf{Y} = \mathbf{s}_1 + \mathbf{N}$$

$$\vdots : \vdots$$

$$H_M : \mathbf{Y} = \mathbf{s}_M + \mathbf{N}$$

where **Y**, **s**_{*i*} and **N** are the projections of y(t), $s_i(t)$ and n(t) respectively onto the signal space spanned by $\{s_i(t)\}$.

Projection of Signals onto Signal Space

- Consider an orthonormal basis {ψ_i(t) | i = 1,...,K} for the space spanned by {s_i(t) | i = 1,...,M}
- Projection of *s_i*(*t*) onto the signal space is

$$\mathbf{S}_i = \begin{bmatrix} \langle \mathbf{S}_i, \psi_1 \rangle & \cdots & \langle \mathbf{S}_i, \psi_K \rangle \end{bmatrix}^T$$



Projection of Observed Signal onto Signal Space

• Projection of y(t) onto the signal space is

$$\mathbf{Y} = \begin{bmatrix} \langle \boldsymbol{y}, \psi_1 \rangle & \cdots & \langle \boldsymbol{y}, \psi_K \rangle \end{bmatrix}^T$$



Projection of Noise onto Signal Space

• Projection of n(t) onto the signal space is

$$\mathbf{N} = \begin{bmatrix} \langle \boldsymbol{n}, \psi_1 \rangle & \cdots & \langle \boldsymbol{n}, \psi_K \rangle \end{bmatrix}^{\mathsf{T}}$$



Proof of Theorem

•
$$\mathbf{Y} = \begin{bmatrix} \langle \mathbf{y}, \psi_1 \rangle & \cdots & \langle \mathbf{y}, \psi_K \rangle \end{bmatrix}^T$$

Component of y(t) orthogonal to the signal space is

$$\mathbf{y}^{\perp}(t) = \mathbf{y}(t) - \sum_{i=1}^{K} \langle \mathbf{y}, \psi_i \rangle \psi_i(t)$$

• y(t) is equivalent to $(\mathbf{Y}, y^{\perp}(t))$

у

• We claim that $y^{\perp}(t)$ is an irrelevant statistic

$$\overset{\perp}{}(t) = \mathbf{y}(t) - \sum_{j=1}^{K} \langle \mathbf{y}, \psi_j \rangle \psi_j(t)$$

$$= \mathbf{s}_i(t) + \mathbf{n}(t) - \sum_{j=1}^{K} \langle \mathbf{s}_i + \mathbf{n}, \psi_j \rangle \psi_j(t)$$

$$= \mathbf{n}(t) - \sum_{j=1}^{K} \langle \mathbf{n}, \psi_j \rangle \psi_j(t) = \mathbf{n}^{\perp}(t)$$

where $n^{\perp}(t)$ is the component of n(t) orthogonal to the signal space.

 n[⊥](t) does not depend on which s_i(t) was transmitted and is independent of N, which makes y[⊥](t) an irrelevant statistic.

M-ary Signaling in AWGN Channel

• *M* hypotheses with prior probabilities π_i , i = 1, ..., M

$$\begin{array}{rcl} H_1 & : & \mathbf{Y} = \mathbf{s}_1 + \mathbf{N} \\ \vdots & & \vdots \\ H_M & : & \mathbf{Y} = \mathbf{s}_M + \mathbf{N} \end{array}$$

• $\mathbf{N} \sim N(\mathbf{m}, \mathbf{C})$ where $\mathbf{m} = \mathbf{0}$ and $\mathbf{C} = \sigma^2 \mathbf{I}$

$$\operatorname{cov}\left(\langle n,\psi_1\rangle,\langle n,\psi_2\rangle\right)=\sigma^2\langle\psi_1,\psi_2\rangle.$$

Optimal Receiver for the AWGN Channel

Theorem (MPE Decision Rule)

The MPE decision rule for M-ary signaling in AWGN channel is given by

$$\begin{split} \delta_{MPE}(\mathbf{y}) &= \underset{1 \leq i \leq M}{\operatorname{argmin}} \|\mathbf{y} - \mathbf{s}_i\|^2 - 2\sigma^2 \log \pi_i \\ &= \underset{1 \leq i \leq M}{\operatorname{argmax}} \langle \mathbf{y}, \mathbf{s}_i \rangle - \frac{\|\mathbf{s}_i\|^2}{2} + \sigma^2 \log \pi_i \end{split}$$

Proof

$$\begin{split} \delta_{MPE}(\mathbf{y}) &= \operatorname*{argmax}_{1 \leq i \leq M} \pi_i p_i(\mathbf{y}) \\ &= \operatorname{argmax}_{1 \leq i \leq M} \pi_i \exp\left(-\frac{\|\mathbf{y} - \mathbf{s}_i\|^2}{2\sigma^2}\right) \end{split}$$

MPE Decision Rule



Continuous-Time Version of MPE Rule

Discrete-time version

$$\delta_{MPE}(\mathbf{y}) = \underset{1 \leq i \leq M}{\operatorname{argmax}} \langle \mathbf{y}, \mathbf{s}_i \rangle - \frac{\|\mathbf{s}_i\|^2}{2} + \sigma^2 \log \pi_i$$

Continuous-time version

$$\delta_{MPE}(\mathbf{y}) = \operatorname*{argmax}_{1 \leq i \leq M} \langle \mathbf{y}, \mathbf{s}_i \rangle - \frac{\|\mathbf{s}_i\|^2}{2} + \sigma^2 \log \pi_i$$

MPE Decision Rule Example



ML Receiver for the AWGN Channel

Theorem (ML Decision Rule)

The ML decision rule for M-ary signaling in AWGN channel is given by

$$\begin{split} \delta_{ML}(\mathbf{y}) &= \underset{1 \leq i \leq M}{\operatorname{argmin}} \|\mathbf{y} - \mathbf{s}_i\|^2 \\ &= \underset{1 \leq i \leq M}{\operatorname{argmax}} \langle \mathbf{y}, \mathbf{s}_i \rangle - \frac{\|\mathbf{s}_i\|^2}{2} \end{split}$$

Proof

$$\begin{split} \delta_{ML}(\mathbf{y}) &= \operatorname*{argmax}_{1 \leq i \leq M} p_i(\mathbf{y}) \\ &= \operatorname{argmax}_{1 \leq i \leq M} \exp\left(-\frac{\|\mathbf{y} - \mathbf{s}_i\|^2}{2\sigma^2}\right) \end{split}$$

ML Decision Rule



ML Decision Rule



Continuous-Time Version of ML Rule

Discrete-time version

$$\delta_{ML}(\mathbf{y}) = \operatorname*{argmax}_{1 \leq i \leq M} \langle \mathbf{y}, \mathbf{s}_i \rangle - \frac{\|\mathbf{s}_i\|^2}{2}$$

Continuous-time version

$$\delta_{ML}(\mathbf{y}) = \underset{1 \leq i \leq M}{\operatorname{argmax}} \langle \mathbf{y}, \mathbf{s}_i \rangle - \frac{\|\mathbf{s}_i\|^2}{2}$$

ML Decision Rule for Antipodal Signaling



where $s_{MF}(t) = s_1(T - t)$ is the matched filter.

References

• Sections 3.3, 3.4, *Fundamentals of Digital Communication*, Upamanyu Madhow, 2008